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**Discussion of „Density Combination at the Norges
Bank“**

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**Discussion of “There is more than one weight to skin a
cat: Combining densities at Norges Bank”
by Bjornland, Gerdrup, Jore, Smith and Thorsrud**

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Outline

- This paper reports an extensive investigation of different ways of constructing combined forecasts, considering both point forecasts and density forecasts for two Norwegian series
- It is found that you do better according to a particular performance measure to use combination weights based on that measure
- I first make some observations of a technical nature, in the course of a presentation of some of the statistical background to such an exercise
- I then say a couple of things I would have done differently

Combining point forecasts (Bates and Granger, 1969; Smith and Wallis, 2009)

Consider two point forecasts of y_t made h periods earlier, f_{1t} and f_{2t}

Forecast errors $e_{it} = y_t - f_{it}$ have variances σ_i^2 and covariance σ_{12}

Combined forecast: $f_{Ct} = kf_{1t} + (1-k)f_{2t}$

Optimal weight: $k_{opt} = \frac{\sigma_2^2 - \sigma_{12}}{\sigma_1^2 + \sigma_2^2 - 2\sigma_{12}}$

Combination using k_{opt} has expected squared error $\sigma_{opt}^2 \leq \min(\sigma_1^2, \sigma_2^2)$

Other weights: “inverse MSE” $k' = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2} = \frac{\frac{1}{\sigma_1^2}}{\frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2}}$; simple average $k = \frac{1}{2}$

Estimated weights: $\text{asy var}(\hat{k}') = (1 - \rho)^2 \text{asy var}(\hat{k}_{opt})$ where $\rho = \sigma_{12} / \sigma_1 \sigma_2$

Combining density forecasts

(1) Linear opinion pool/ finite mixture distribution (Wallis, 2005)

Consider n individual density forecasts $f_i(y)$, $i = 1, \dots, n$, for given t and h

Combined forecast: $f_C(y) = \sum_{i=1}^n w_i f_i(y)$ with weights $w_i \geq 0$, $i = 1, \dots, n$, $\sum w_i = 1$

If $f_i(y) = N(\mu_i, \sigma_i^2)$ then $f_C(y)$ is a mixture of normals (Pearson, 1894)

For any individual densities, moments about the origin are mixed in the same way:

$$\mu'_1 = \sum_{i=1}^n w_i \mu_i = \mu_C, \quad \mu'_2 = \sum_{i=1}^n w_i (\mu_i^2 + \sigma_i^2)$$

hence the variance of f_C is $\sigma_C^2 = \mu'_2 - \mu_C^2 = \sum_{i=1}^n w_i \sigma_i^2 + \sum_{i=1}^n w_i (\mu_i - \mu_C)^2$

i.e. aggregate uncertainty = average individual uncertainty + disagreement

(2) Logarithmic combination (Bacharach, Hammond)

This is usually written as $f_C(y) = \frac{\prod f_i^{w_i}(y)}{\int \prod f_i^{w_i}(y) dy}$

If $f_i(y) = N(\mu_i, \sigma_i^2)$ then $f_C(y) = N(\mu_C, \sigma_C^2)$ where

$$\frac{\mu_C}{\sigma_C^2} = \sum_{i=1}^n w_i \frac{\mu_i}{\sigma_i^2} \quad \text{and} \quad \frac{1}{\sigma_C^2} = \sum_{i=1}^n w_i \frac{1}{\sigma_i^2}$$

In the case of equal weights, σ_C^2 is the harmonic mean of the individual variances; this is less than the arithmetic mean, which is less than the finite mixture variance

(3) Pragmatic combination (Wallis, 2005)

Combine moments as in the finite mixture distribution, and use a normal (or 2PN) distribution with these moments to produce interval forecasts and fan charts

Evaluating combined density forecasts

- For density forecasts we have no results like the Bates-Granger result for the **expected** squared error of the optimal point forecast combination
- For several scoring rules we have the result that the **expected** score of the combined density forecast is between that of the best and worst forecasts
- This suggests removing (“trimming”) the worst forecasts from the collection before constructing combined forecasts
 - trim the set of forecasting models
 - for point forecasting, remove extremes from the set $\{f_{it}\}$
- We have individual examples where the **actual** score of the “best” combination is better than that of the best forecast (in “Optimal prediction pools” by John Geweke and Gianni Amisano, October 2008)

A public health warning about forecast evaluation criteria

Some criteria are not sensitive to the use of the wrong information set

- point forecasts: tests of equality of in-sample and out-of-sample MSE
- density forecasts: tests of uniformity of probability integral transforms
 - any forecast that is the correct conditional distribution for the information on which it is based has uniform PITs
 - extreme example: the unconditional distribution

So we need additional criteria, and **comparative** testing (“encompassing”)

Some things I would do differently

- give a clear account of the statistical framework and definitions of all statistics employed
- make more use of the estimation sample results to rule out rubbish models
- separate the discussion of the point and density forecasting problems, since they are different problems; only then take a brief look at similarities and differences between the rankings of models/methods in the two problems
- give plots of the target series and selected competing forecasts, so that the practical significance of perceived differences can be assessed
- use more target series
- in computing inverse MSE weights for use in constructing forecast combinations, consider Stock and Watson's variations:
 - raise each $1/\text{MSE}$ term to the power ω : $\omega < 1$ shrinks weights towards equality, $\omega > 1$ increases the weight given to better-performing forecasts
 - increase weight on recent errors via discounted sum of squared errors